# Written Exam for the M.Sc. in Economics Autumn 2011-2012 (Fall Term) 

Financial Econometrics A: Volatility Modelling

Final Exam: Masters course
Exam date: 11/1-2012

## 3-hour open book exam.

Please note all questions need to be answered.
Please note that the language used in your exam paper must correspond to the language of the title for which you registered during exam registration. I.e. if you registered for the English title of the course, you must write your exam paper in English. Likewise, if you registered for the Danish title of the course or if you registered for the English title which was followed by "eksamen på dansk" in brackets, you must write your exam paper in Danish.

If you are in doubt about which title you registered for, please see the print of your exam registration from the students' self-service system.

## Question A:

Question A.1: Figure A. 1 shows a log-returns series $y_{t}$ with $T=1000$ observations together with the ACF for $\left|x_{t}\right|$. Furthermore, Table A. 1 shows some output from estimation of a $\operatorname{GARCH}(1,1)$ model with these data.

From Figure A.1, would you expect a $\operatorname{GARCH}(1,1)$ to fit the data? Explain the output in Table A.1.


Figure A. 1

| Table A.1 |  |
| :--- | :--- |
| Parameter estimates in GARCH $(1,1):$ | $\hat{\alpha}+\hat{\beta}=1.03$ |
| Standardized residuals: $\hat{z}_{t}=y_{t} / \hat{\sigma}_{t}$ |  |
| Normality Test for $\hat{z}_{t}:$ | p-value: 0.00 |
| LM ARCH test in $\hat{z}_{t}:$ | p-value: 0.15 |

Question A.2: Theory suggests that the volatility of $y_{t}$ is driven by the an observed factor $x_{t}$, and not by its own past as suggested by estimating the $\operatorname{GARCH}(1,1)$ model. We wish to investigate if indeed $y_{t}$ 's conditional volatility $\sigma_{t}$ is driven by $x_{t}$ and do so by considering the following model:

$$
\begin{align*}
y_{t} & =\sigma_{t} z_{t}  \tag{1}\\
\sigma_{t}^{2} & =\omega+\alpha_{x} x_{t-1}^{2}+\alpha_{y} y_{t-1}^{2} \tag{2}
\end{align*}
$$

where $z_{t}$ are $\operatorname{iidN}(0,1)$ and $x_{t}$ is the observed process which drives the conditional variance of the returns $y_{t}$. The parameters to be estimated are given by $\theta=\left(\omega, \alpha_{x}, \alpha_{y}\right)$, with $\omega>0$ and $\alpha_{x}, \alpha_{y} \geq 0$.

State the log-likelihood function $L_{T}(\theta)$ and show that $S_{T}(\theta):=\frac{1}{\sqrt{T}} \partial L_{T}(\theta) / \partial \alpha_{x}$, is given by (up to your choice of normalization),

$$
S_{T}(\theta)=\frac{1}{\sqrt{T}} \sum_{t=1}^{T} \frac{1}{2}\left(\frac{y_{t}^{2}}{\sigma_{t}^{2}}-1\right)\left(\frac{x_{t-1}^{2}}{\sigma_{t}^{2}}\right) .
$$

Question A.3: Show that $S_{T}\left(\theta_{0}\right)$ is asymptotically Gaussian. When you establish this, explain which results you use, and how you use them. Also, state and discuss which conditions on the parameters $\theta_{0}$ and on the exogenous process $x_{t}$ you use.

Discuss briefly what is missing in order to conclude that $\sqrt{T}\left(\hat{\alpha}_{x}-\alpha_{x, 0}\right)$ is asymptotically Gaussian.

## Question A.4:

Suppose that the factor $x_{t}$ driving $\sigma_{t}$ is given by an $\operatorname{AR}(1)$ process,

$$
\begin{equation*}
x_{t}=\rho x_{t-1}+\varepsilon_{t}, \tag{3}
\end{equation*}
$$

where $\varepsilon_{t}$ are $\operatorname{iidN}\left(0, \sigma_{\varepsilon}^{2}\right)$ and the two innovation series $\varepsilon_{t}$ and $z_{t}$ are independent. The series $x_{t}$ and its differences $\Delta x_{t}$ are shown in Figure A.2. Comment on the output reported in Table A. 2 using Figure A. 2 and your previous results.

| Table A.2 |  |
| :---: | :--- |
| Model: | Estimates/[std. error] |
| $\sigma_{t}^{2}=\omega+\alpha_{x} x_{t}^{2}+\alpha_{y} y_{t-1}^{2}$ | $\hat{\alpha}_{x}=\underset{[0.4}{0.4} \hat{\alpha}_{y}=0.03$ |
| Standardized residuals: $\hat{z}_{t}=y_{t} / \hat{\sigma}_{t}$ |  |
| $029]$ |  |
| Normality Test for $\hat{z}_{t}:$ | p-value: 0.72 |
| LM ARCH test in $\hat{z}_{t}:$ | p-value: 0.85 |



Figure A. 2

## Question B:

Question B.1: Consider the return series $y_{t}$ in Figure B. 1 with $\mathrm{t}=1,2, \ldots, \mathrm{~T}=3000$.


Figure B. 1
Estimation with a 2-state Markov switching volatility model, gave the following output in the usual notation in terms of the transition matrix $P=$ $\left(p_{i j}\right)_{i, j=1,2}$ and smoothed standardized residuals $\hat{z}_{t}^{*}$ :

| $\hat{P}$, QMLE of $P:$ | $\hat{p}_{11}=0.98 \quad \hat{p}_{21}=0.044$ |
| :--- | :--- |
| Normality test for $\hat{z}_{t}:$ | p-value: 0.000 |
| LM ARCH test in $\hat{z}_{t}:$ | p-value: 0.000 |

What would you conclude on the basis of the output and the graph?
Question B.2: As an alternative to the 2-state Markov switching volatility model consider instead the 3 -state switching model for $\log$-returns $y_{t}$ given by,

$$
y_{t}=\sigma_{s_{t}} z_{t}
$$

with $z_{t} \operatorname{iidN}(0,1)$ and $s_{t}=1,2$ or 3 such $\sigma_{s t}$ can take three values, $\sigma_{1}, \sigma_{2}$ and $\sigma_{3}$ ("low", "medium", "high" say) according to the value of the switching variable $s_{t}$.

Initially, we let the switching variable $s_{t}$ be iid with $P\left(s_{t}=1\right)=p_{1}$, $P\left(s_{t}=2\right)=p_{2}$ and finally $P\left(s_{t}=3\right)=p_{3}=1-p_{1}-p_{2}$.

Discuss if $\sigma_{s_{t}}^{2}$ and $y_{t}$ are weakly mixing and stationary if the three probabilities $p_{1}, p_{2}$ and $p_{3}$ satisfy $0<p_{i}<1$.

Explain how you would write an algorithm in for example ox which can simulate the $s_{t}$ series given fixed values of $p_{1}$ and $p_{2}$. You do not need to write the actual ox code, but should explain how you simulate $s_{t}$ by drawing suitable iid series.

Question B.3: Given return data $Y:=\left(y_{t}\right)_{t=1,2,3, \ldots, T}$ we wish to estimate the parameters in $\theta=\left\{\sigma_{1}^{2}, \sigma_{2}^{2}, \sigma_{3}^{2}, p_{1}, p_{2}\right\}$.

Let $L(Y, S ; \theta)$ be the likelihood function with both $Y$ and $S:=\left(s_{t}\right)_{t=1,2, \ldots, T}$ observed. Find $L_{E M}(Y ; \theta)$ where,

$$
L_{E M}(Y ; \theta)=E_{\tilde{\theta}}(\log L(Y, S ; \theta) \mid Y) .
$$

Explain in particular how the so-called smoothed probabilities $\left(p_{i t}^{*}\right)_{i=1,2,3}$ enter in $L_{E M}(Y ; \theta)$.

Question B.4: Show that maximizing $L_{E M}(Y ; \theta)$ one gets,

$$
\hat{\sigma}_{1}^{2}=\frac{\sum_{t=1}^{T} p_{1, t}^{*} y_{t}^{2}}{\sum_{t=1}^{T} p_{1, t}^{*}} .
$$

Give an interpretation of $\hat{\sigma}_{1}^{2}$. Also explain how $p_{1 t}^{*}$ is computed.
Question B.5: It is found that a better description for the data at hand is given by letting $s_{t}$ be Markov switching according to the transition matrix,

$$
P=\left(\begin{array}{lll}
p_{11} & p_{21} & p_{31}  \tag{4}\\
p_{12} & p_{22} & p_{32} \\
p_{13} & p_{23} & p_{33}
\end{array}\right),
$$

where $\sum_{i=1}^{3} p_{1 i}=\sum_{i=1}^{3} p_{2 i}=\sum_{i=1}^{3} p_{3 i}=1$. It is empirically found that $H_{0}: \sigma_{2}=\sigma_{3}=\sigma^{*}$ holds.

Interpret and comment on $H_{0}$.
Find under the assumption that $H_{0}$ holds, the two probabilities:

$$
P\left(\sigma_{t}=\sigma^{*} \mid \sigma_{t-1}=\sigma_{1}\right) \quad \text { and } \quad P\left(\sigma_{t}=\sigma^{*} \mid \sigma_{t-2}=\sigma_{1}\right) .
$$

